

Sample size Re-estimation and Bias

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Content

- Sample size re-estimation, bias and correction
- Continuous variance monitoring
- Blinded continuous variance monitoring
- More insight into the bias and it's consequences

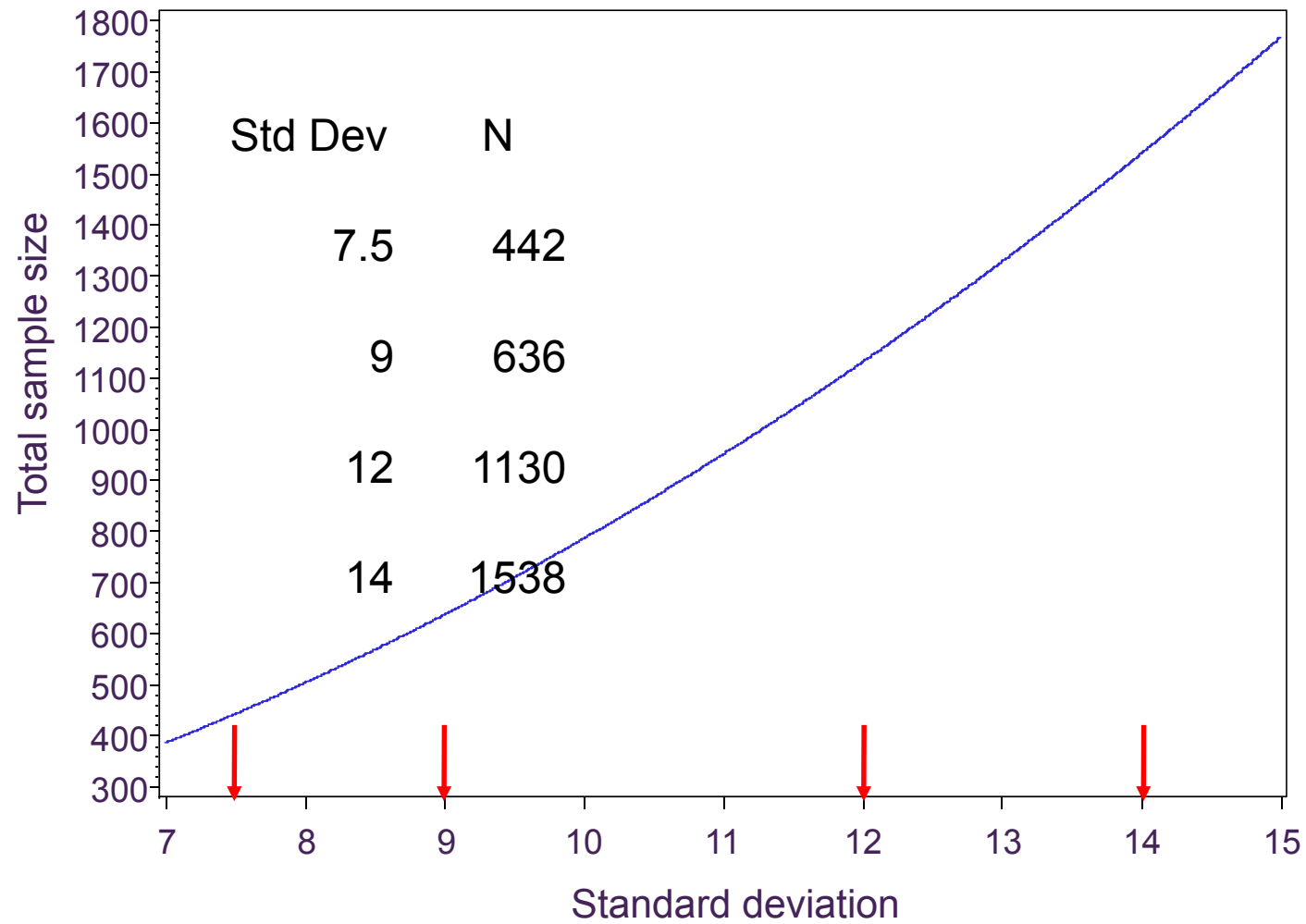
Clinical study and sample size calculation

- Randomized study comparing two treatments (e.g. new treatment versus control)
- A variable is observed which is assumed to be independently **normally distributed** with common unknown variance σ^2 and treatment difference δ
- Significance test with level α and power $1-\beta$ for $\delta = \delta_a$
 $H_0 : \delta = 0$ (new treatment equal to control) versus
 $H_a : \delta > 0$ (new treatment better than control)
- Sample size per treatment $n = v\sigma^2$, $v = 2(z_{1-\alpha} + z_{1-\beta})^2 / \delta_a^2$
- Since σ^2 is unknown, an initial guess required

Example: MacDonald et al. (2008)

- **Objective:** Assessment Lumiracoxib's effect on blood pressure in Osteoarthritis patients with hypertension
- **Treatments:** Lumiracoxib or Ibuprofen
- **Primary endpoint:** change from baseline at week 4 in average 24 h systolic blood pressure
- Significance level $\alpha=0.025$ (1-sided), power 80% for $\delta_a = 2$ mmHg
- **Standard deviation $\sigma = ???$ mmHg**
 - White et al (2002): **9** mmHg observed (but different population)
 - Sowers et al (2005): assumed **7.5**, observed **12** mmHg (6 week follow-up)
 - Other studies in non-OA population: up to **14** mmHg

Uncertainty in the planning phase



Uncertainty about variance

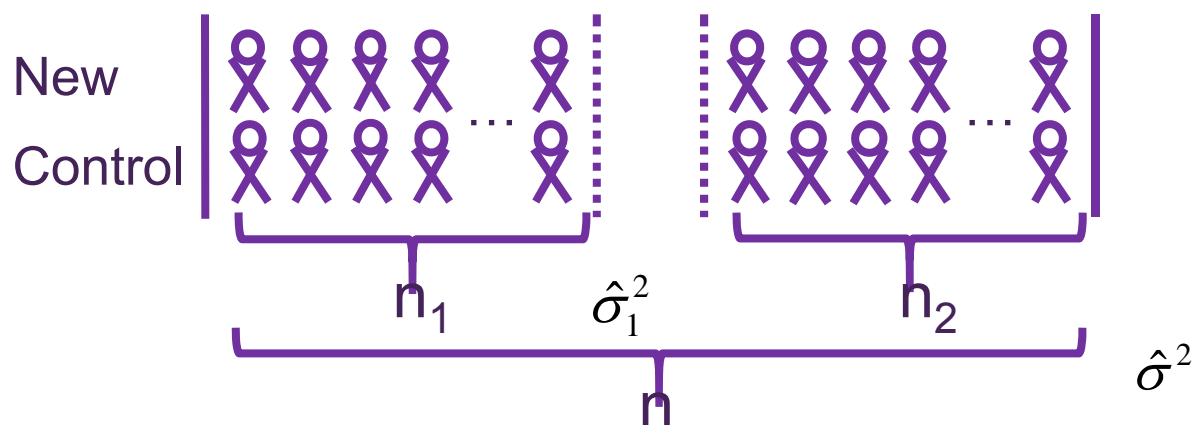
- Problem: Often considerable uncertainty regarding the variance
- Solutions include:
 - **One interim look:** estimate σ^2 from accruing data in an interim sample size review and use this estimate to adjust sample size
 - **Several interim looks** with update of variance estimate and sample size
 - **Continuous monitoring** of the variance

Sample size based on one interim analysis

Observe n_1 patients ("Stage 1") & estimate variance: $\hat{\sigma}_1^2$

Final sample size $n = v\hat{\sigma}_1^2$

Observe additional $n_2 = n - n_1$ patients ("Stage 2")

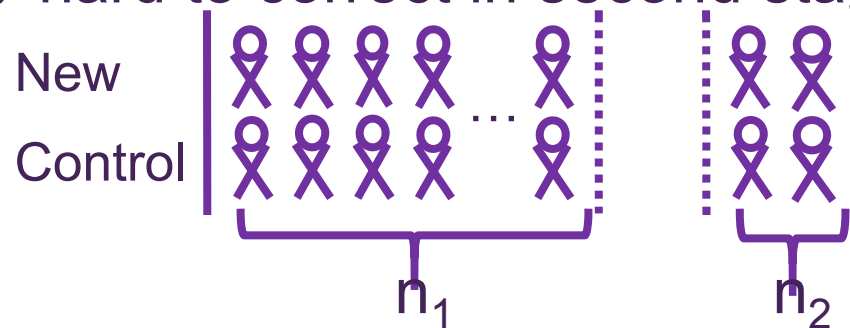


- The variance estimator $\hat{\sigma}^2$ at the end of the trial is biased. It underestimates the true variance!

Why is there a bias?

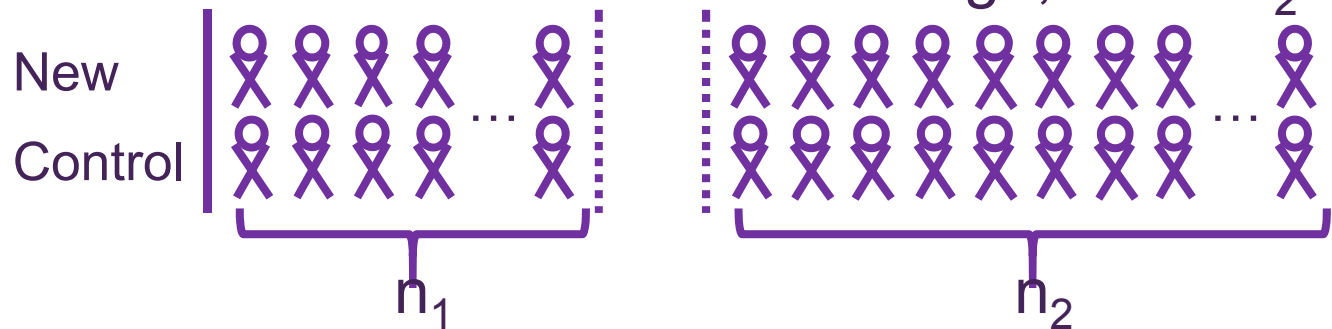
$\hat{\sigma}_1^2$ lower than true σ^2

→ hard to correct in second stage, since n_2 low



• $\hat{\sigma}_1^2$ larger than true σ^2

→ easier to correct in second stage, since n_2 large



➤ $\hat{\sigma}^2$ underestimates the variance

Consequence of the bias in the final variance estimator

➤ The t-test $\text{Reject } H_0 \Leftrightarrow t > t_{2n-2, 1-\alpha}$ with

$t = \hat{\delta} / \sqrt{2\hat{\sigma}^2 / n}$ does not control the alpha level

How large is the bias?

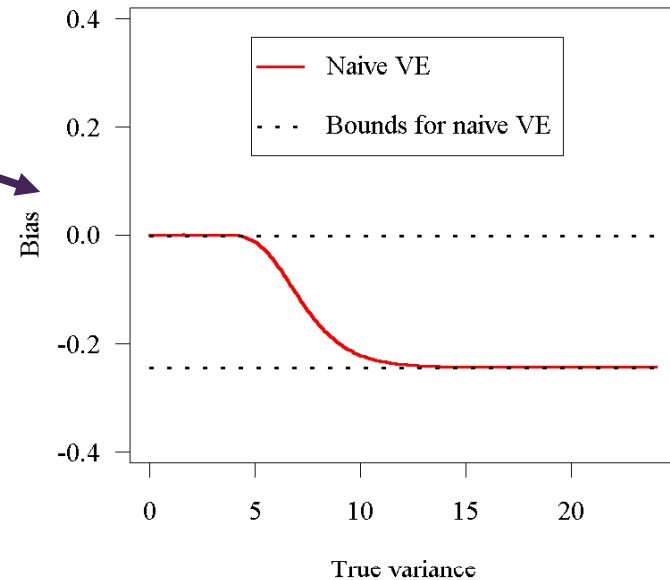
- Bias of the (“naïve”) variance estimator can be computed
- Bounds for bias (Miller, 2005)

$$-\frac{n_1 - 1}{n_1 - 2} \cdot \frac{1}{v} \leq E\hat{\sigma}^2 - \sigma^2 \leq 0$$

recall: $n = v\hat{\sigma}_1^2$, $v = 2(z_{1-\alpha} + z_{1-\beta})^2 / \delta_a^2$

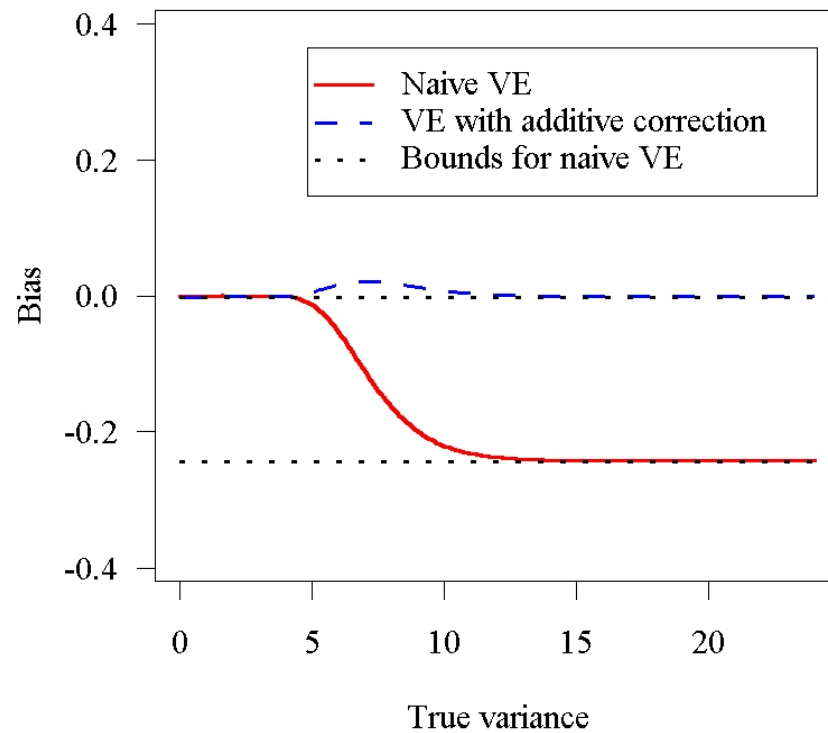
- Additive correction of final variance estimate:

$$\hat{\sigma}_{ac}^2 = \hat{\sigma}^2 + \begin{cases} 0, & \text{if stopped directly after interim} \\ \frac{n_1 - 1}{n_1 - 2} \cdot \frac{1}{v}, & \text{otherwise} \end{cases}$$

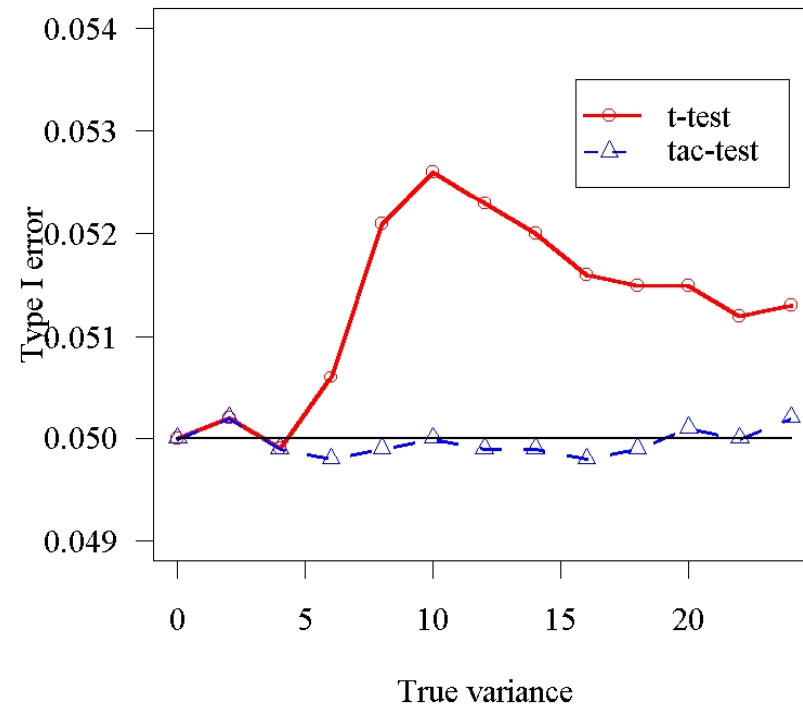


Correction of bias

Bias of variance estimator



Type I error of t-test and t-test with additive correction



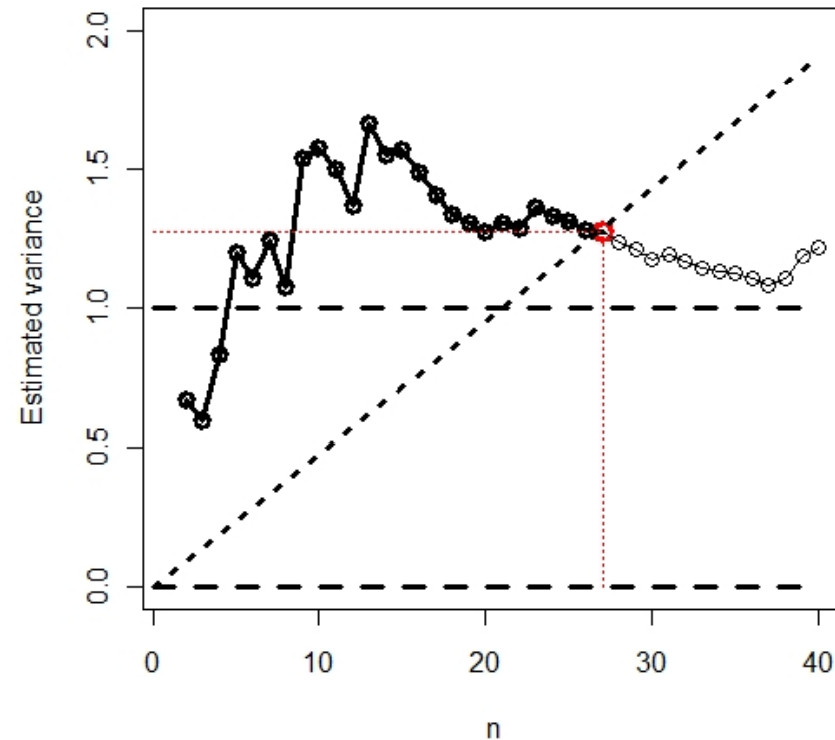
Continuous variance monitoring

- **Continuous variance monitoring procedure:**
 - Monitor variance after each (pair of) patients: $\hat{\sigma}_n^2$ starting after n_1 patients ($n_1 \geq 2$)
 - Stop study as soon as “sample size sufficient” according to this estimate ($n \geq v\hat{\sigma}_n^2$)
- This is a stochastic process with stop-time
$$N = \min \left\{ n = n_1, n_1 + 1, \dots \mid \hat{\sigma}_n^2 \leq n/v \right\}$$
- Discussed in the context of clinical studies e.g. by Mehta & Tsiatis (2001) and Jennison & Turnbull (2007)
- Investigated by Friede & Miller (2012)

Continuous variance monitoring as stochastic process

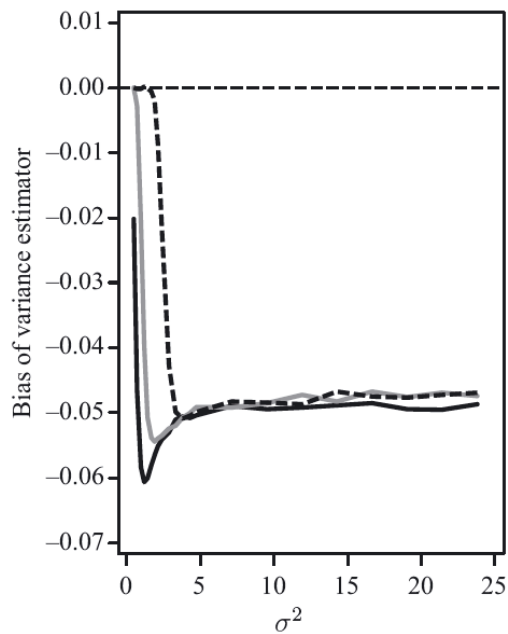
- Example study with true variance=1, $v=21$
$$v = 2(z_{1-\alpha} + z_{1-\beta})^2 / \delta_a^2$$

 $v = 21.0$ for $\alpha = 0.025, \beta = 0.1, \delta_a = 1$
- First time under linear boundary stops study
- $E(\hat{\sigma}_N^2) < \sigma^2 = 1$
(negative bias of variance estimator)

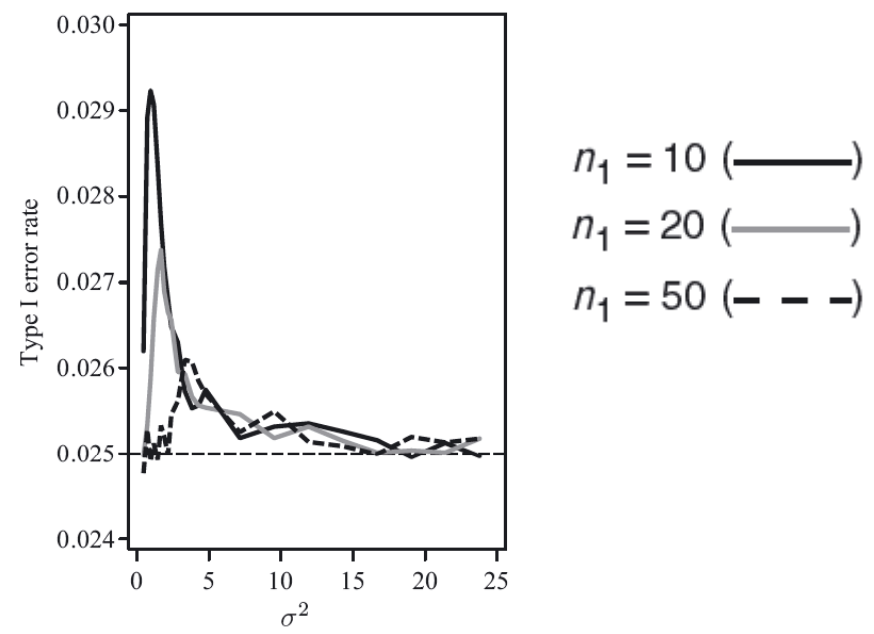


Bias of variance estimator and test

Bias of variance estimator



Bias of test size



Variance estimator $\hat{\sigma}_N^2$ at stop-time N negatively biased

Therefore, the t-test $\text{Reject } H_0 \Leftrightarrow t > t_{2N-2, 1-\alpha}$ with $t = \hat{\delta} / \sqrt{2\hat{\sigma}_N^2 / N}$ does not control the alpha level

Blinding in clinical trials

- **Randomized clinical trials** for drug development are usually **blinded**
- Database of ongoing study has no treatment information
- Separate file with treatment info kept secretly

Database

Patient	Treatment	Result1	Result2...
1	*	1	7.7 ...
2	*	0	6.8 ...
3	*	1	8.0 T
4	*	0	6.5 ...
5	*	0	8.9 ...

Treatments

Patient	Treatment
1	Placebo
2	Active
3	Active
4	Placebo
5	Active

Blinding and sample size re-estimation

- To perform the sample size re-estimation shown before, the **treatment of all patients needs to be known for the computation of the variance (unblinding necessary)**
- Regulatory authorities **prefer methods not requiring unblinding**

$$\hat{\sigma}_n^2 = \frac{1}{2n-2} \left(\sum_{j=1}^n (X_{1j} - \bar{X}_{1\cdot})^2 + \sum_{j=1}^n (X_{2j} - \bar{X}_{2\cdot})^2 \right)$$

↓

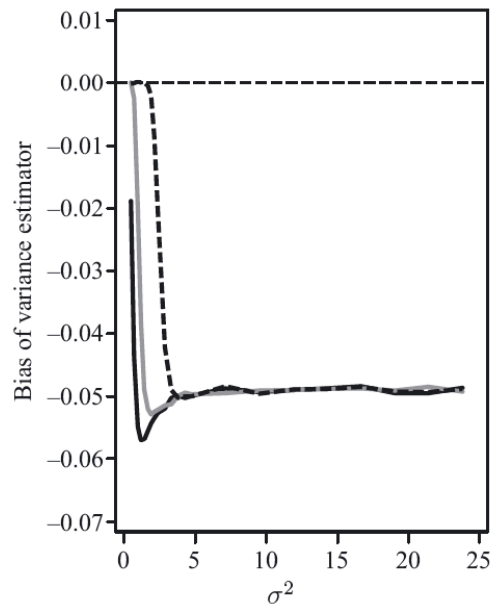
$$\hat{\sigma}_{n,\text{blind}}^2 = \frac{1}{2n-1} \left(\sum_{j=1}^n (X_{1j} - \bar{X}_{\cdot\cdot})^2 + \sum_{j=1}^n (X_{2j} - \bar{X}_{\cdot\cdot})^2 \right)$$

Blinded continuous monitoring

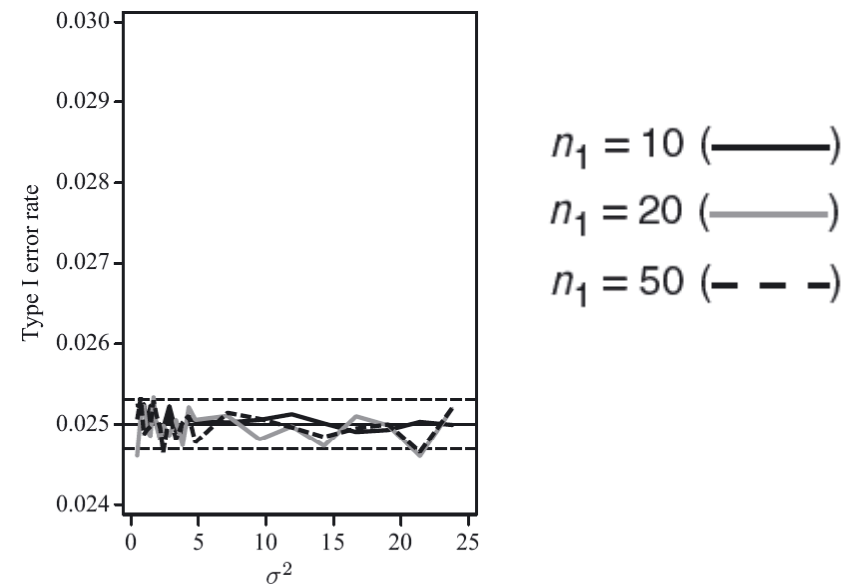
- The different sample size re-estimation procedures can be performed blinded
- Here we show the **blinded version for continuous monitoring**
- Blinded continuous monitoring procedure:
 - Monitor **overall (one sample) variance** $\hat{\sigma}_{n,\text{blind}}^2$ after each (pair of) patients, ignoring the different treatments
 - Stop study as soon as sample size sufficient according to this blinded estimate ($n / \hat{\sigma}_{n,\text{blind}}^2 \geq \nu$)
 - Estimate final variance unblinded

Bias of variance estimator and test – blinded procedure

Bias of variance estimator

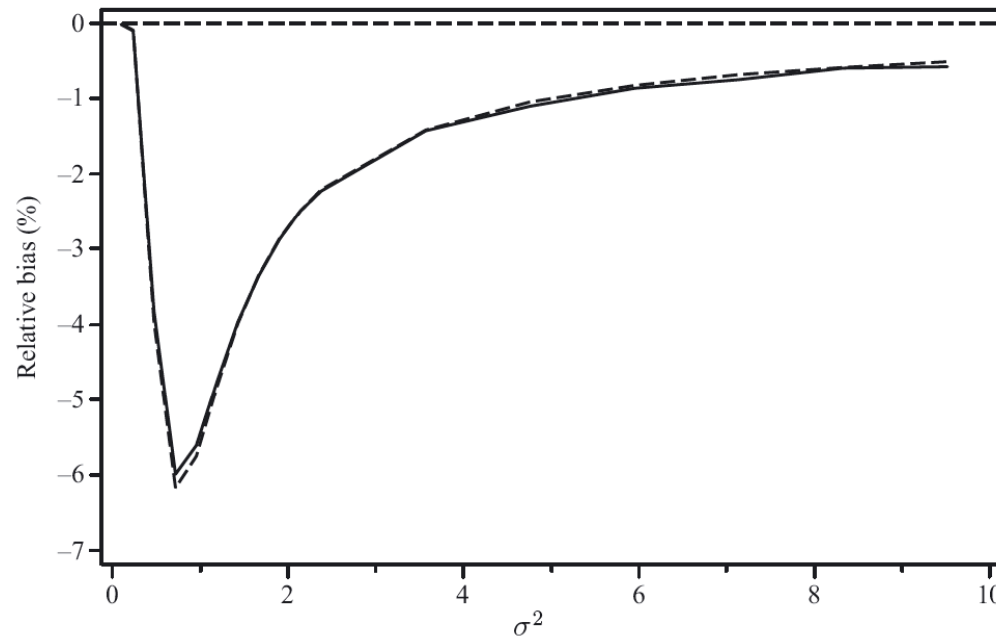


Bias of test size



Still bias for variance estimation but essentially no bias for test.

Blinded procedure: Why is the test size (almost) unbiased?



Relative bias of the numerator (—) and denominator (---) of the F -test

$$t = \hat{\delta}_N / \sqrt{2\hat{\sigma}_N^2 / N}$$

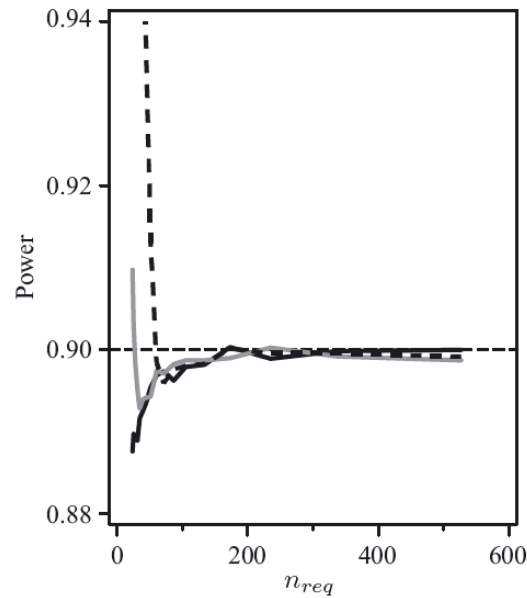
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$$F = \frac{\frac{N}{2} \hat{\delta}_N^2}{\hat{\sigma}_N^2}$$

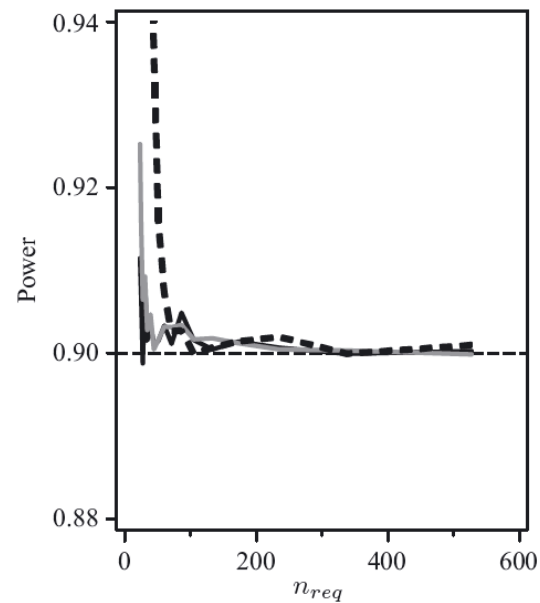
If the (two-sided) test is written as F -test, the **numerator** and the **denominator** have the same bias after blinded continuous monitoring

Power of test after continuous monitoring

Unblinded continuous monitoring



Blinded continuous monitoring

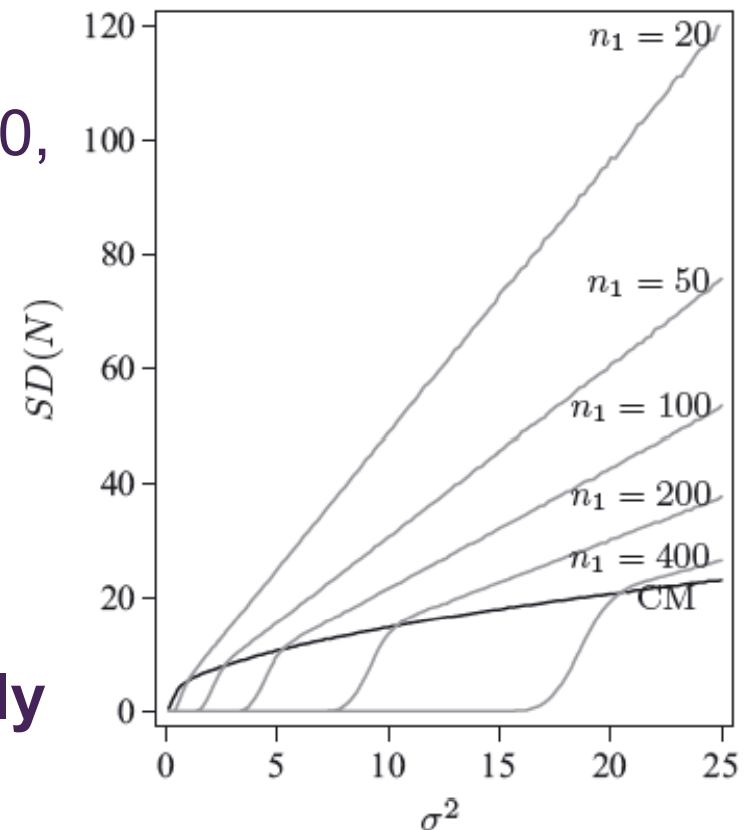


$n_1 = 10$ (—)
 $n_1 = 20$ (---)
 $n_1 = 50$ (· · ·)

Desired power approximately maintained; better so for the blinded procedure.

Blinded procedures: Continuous monitoring versus sample size re-estimation with one interim look

- $\alpha = 0.025, 1-\beta = 0.90, \delta = \delta_a = 1$
- One look after $n_1=20, 50, 100, 200,$ or 400 patients per treatment to estimate variance
- Compare with continuous monitoring (CM)
- If n_1 chosen too large: risk to overshoot necessary sample size
- If n_1 chosen too small: **SD for sample size can be considerably higher than for CM**



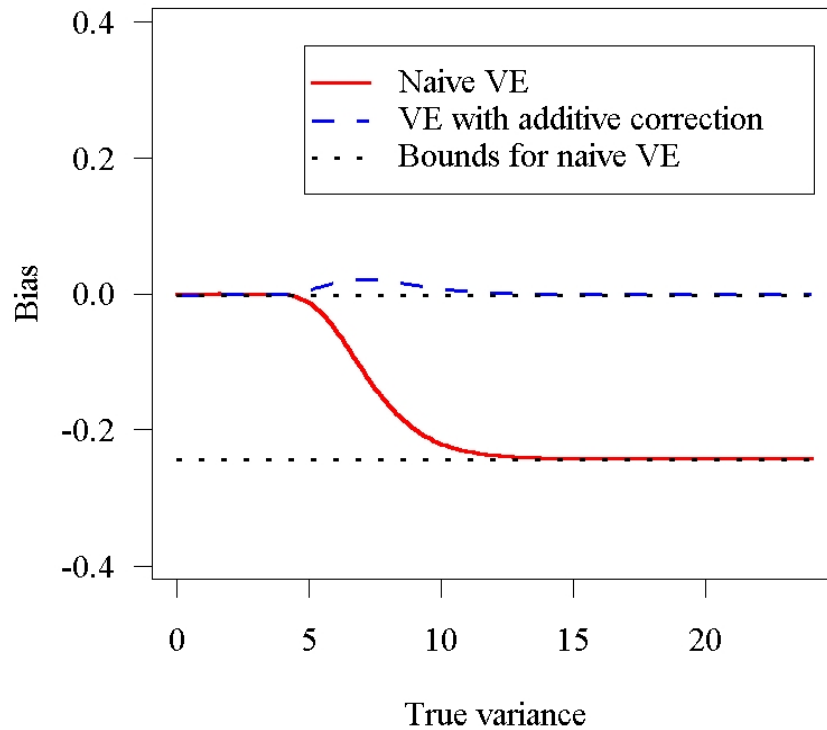
Continuous monitoring – is it logistically feasible?

- Increased use of electronic data capture techniques in clinical studies
- E.g. in chronic pain studies
 - Patients are provided with electronic diaries for their daily pain recording
 - SMS reporting has successfully been applied (Axén, Bodin, Bergström et al, 2012)
 - Pain intensity ratings can immediately be transferred to central database of sponsor
- Continuous monitoring feasible in some (but not all) clinical study situations

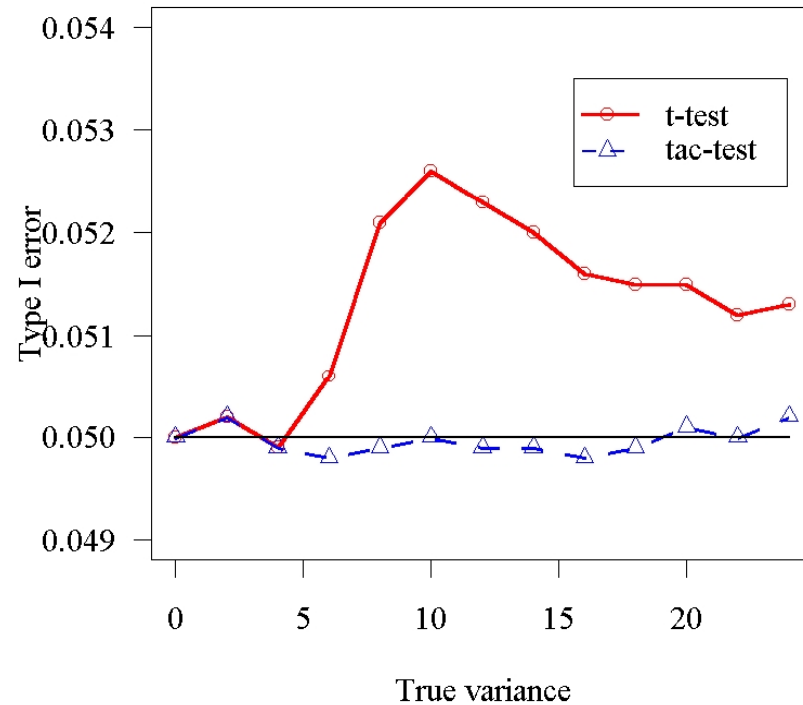


Back to the bias for sample size re-estimation with one look

Bias of variance estimator

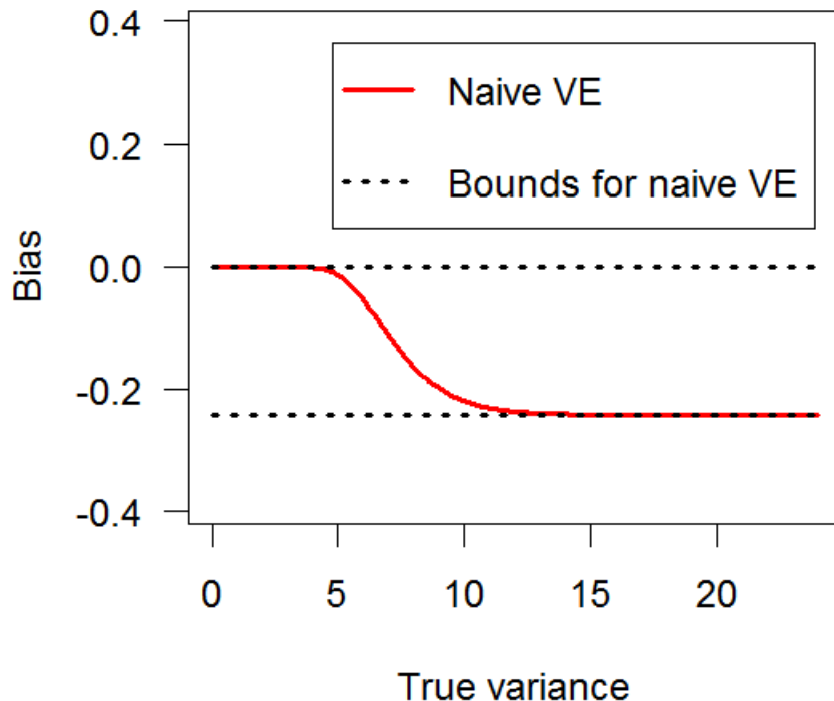


Type I error of t-test and t-test with additive correction



Bias for unblinded case with 1 IA

Bias of variance estimator



$$-\frac{n_1 - 1}{n_1 - 2} \cdot \frac{1}{v} \leq E\hat{\sigma}^2 - \sigma^2 \leq 0$$

Exact bias:

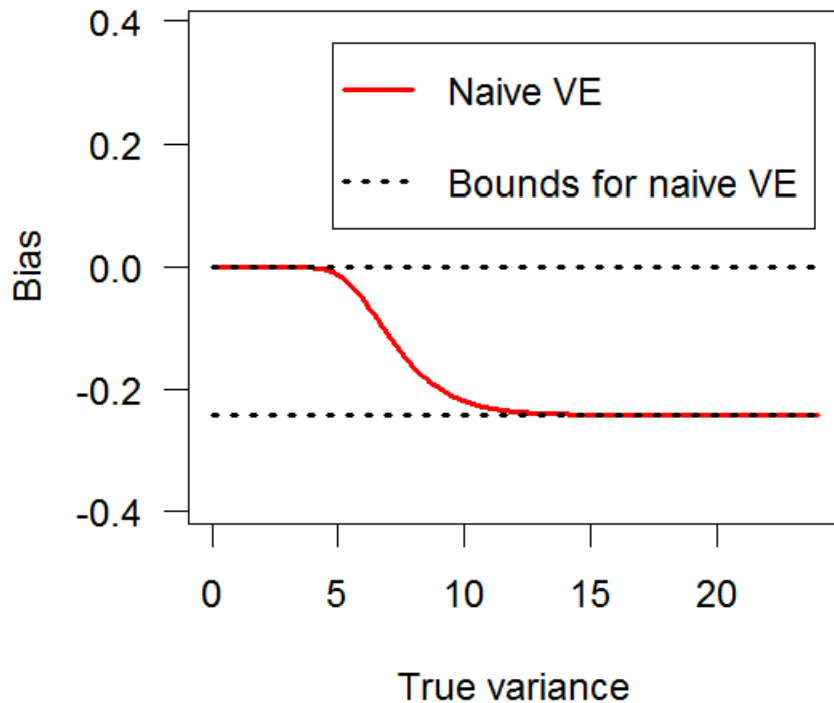
$$\frac{2(n_1 - 1)^2}{vd} \{F_{2n_1}(d) - F_{2n_1-2}(d)\} + \frac{n_1 - 1}{v} \{1 - F_{2n_1-2}(d)\} - \frac{(n_1 - 1)^2}{v(n_1 - 2)} \{1 - F_{2n_1-4}(d)\}.$$

where F is the chi-square distribution function and

$$d = \frac{(2n_1 - 2)(n_1 + n_{2min} - 1)}{v\sigma^2}$$

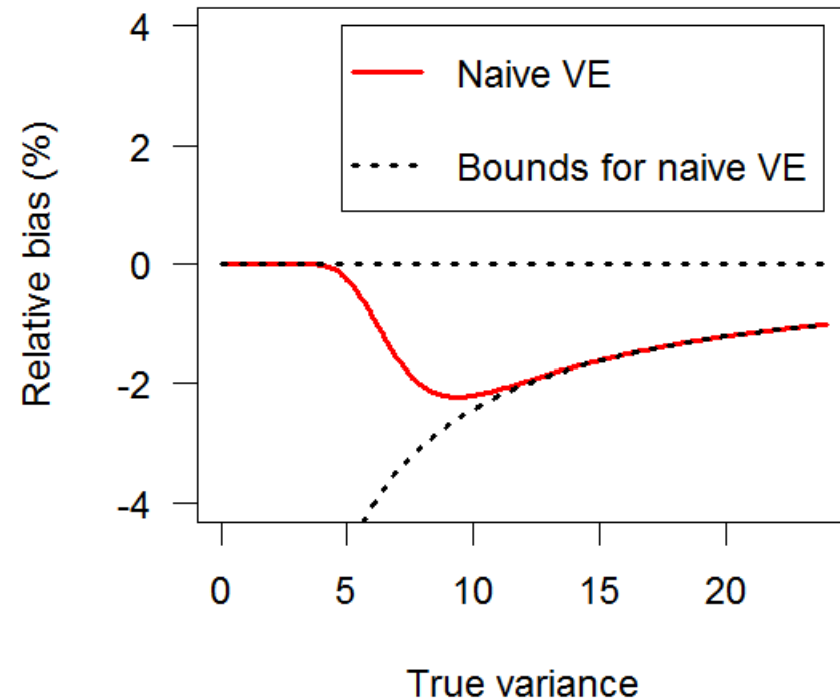
Bias for unblinded case with 1 IA

Absolute bias of variance estimator



$$-\frac{n_1 - 1}{n_1 - 2} \cdot \frac{1}{v} \leq E\hat{\sigma}^2 - \sigma^2 \leq 0$$

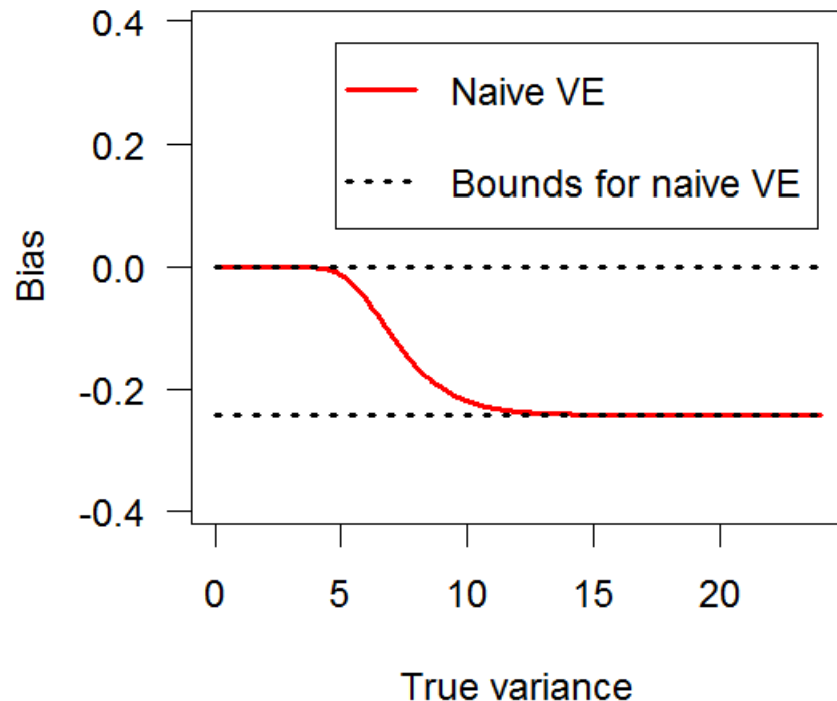
Relative bias of variance estimator



$$-\frac{n_1 - 1}{n_1 - 2} \cdot \frac{1}{v\sigma^2} \leq -\frac{E\hat{\sigma}^2 - \sigma^2}{\sigma^2} \leq 0$$

Bias for unblinded case with 1 IA

Absolute bias of variance estimator



$$-\frac{n_1 - 1}{n_1 - 2} \cdot \frac{1}{v} \leq E\hat{\sigma}^2 - \sigma^2 \leq 0$$

Assumptions for this graph

- Sample size in first stage $n_1 = 20$ per group
- Minimum sample size in second stage: $n_{2\min} = 10$
- $\alpha = 0.025$, power = 90%, $\delta_a = 2.2$

$$v = 2(z_{1-\alpha} + z_{1-\beta})^2 / \delta_a^2$$

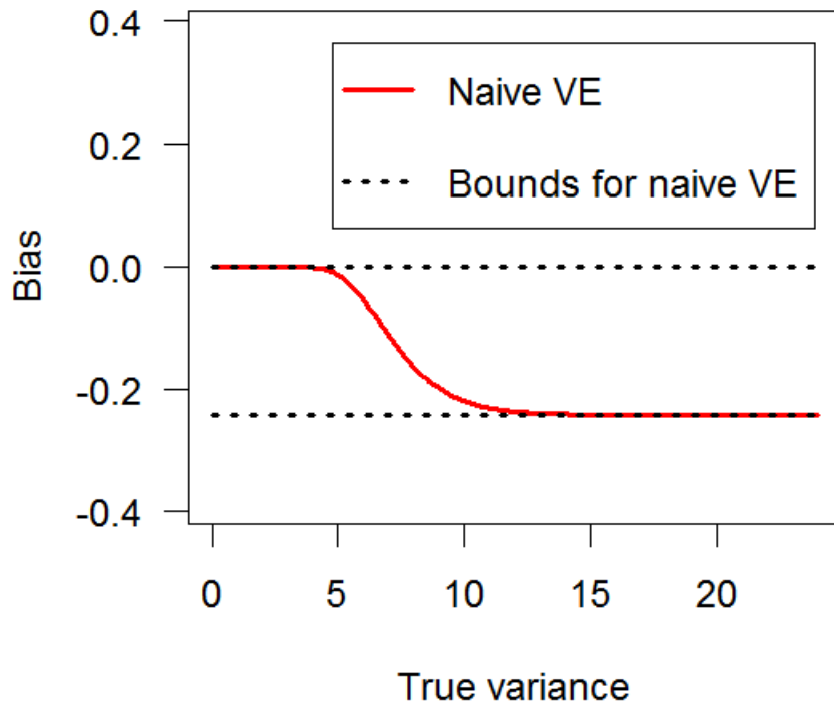
- Sample size formula for interim

$$N = \max \{ v\hat{\sigma}_1^2 + 1, n_1 + n_{2\min} \}$$

- How does the bias depend on $n_{2\min}$, n_1 , δ_a ? \top

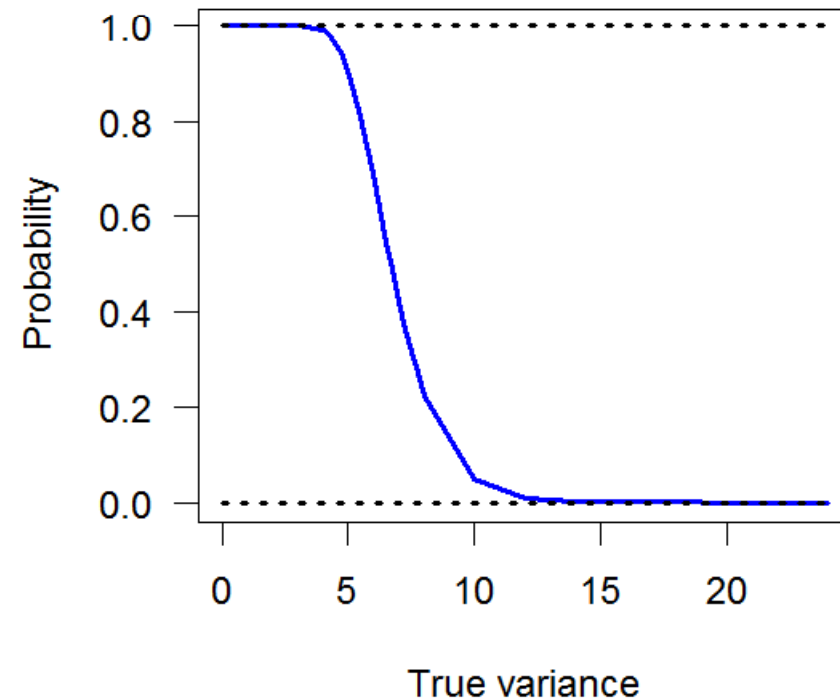
Bias for unblinded case with 1 IA

Absolute bias of variance estimator



$$E\hat{\sigma}^2 - \sigma^2$$

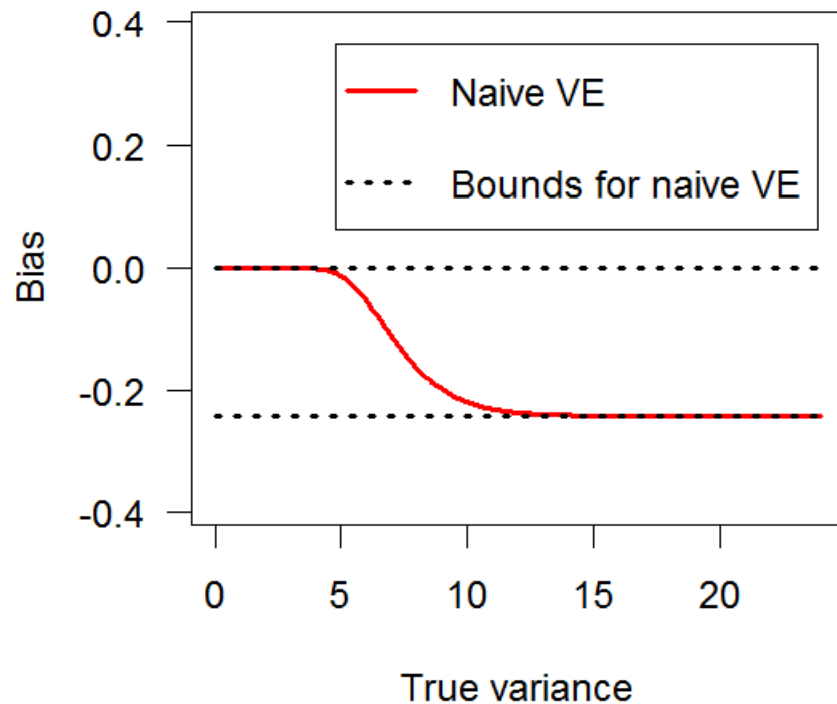
Probability for stopping the study with minimal sample size $n_1 + n_{2\min}$



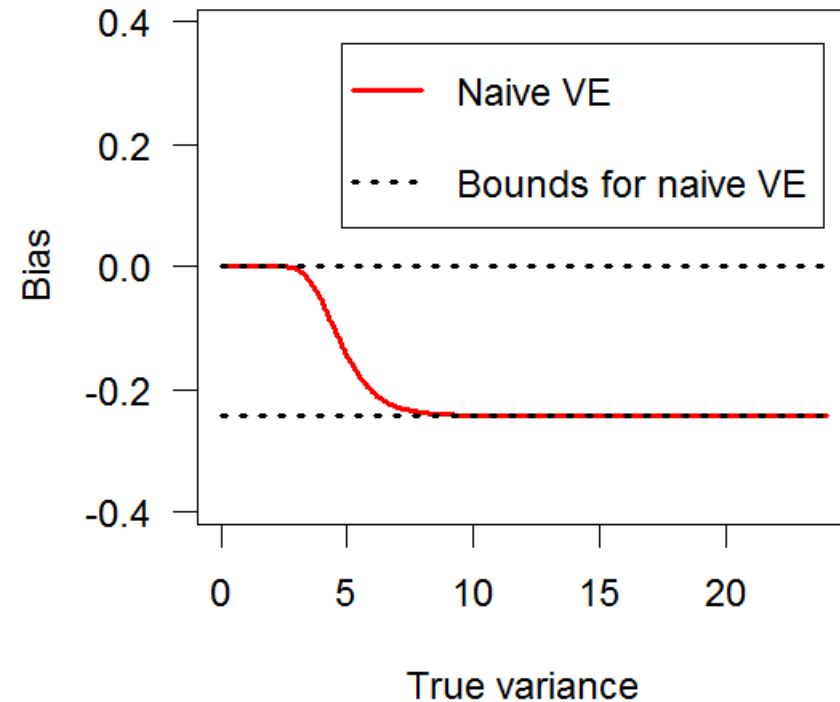
$$P(N = n_1 + n_{2\min})$$

Bias for unblinded case with 1 IA

Bias of variance estimator, $n_{2\min}=10$

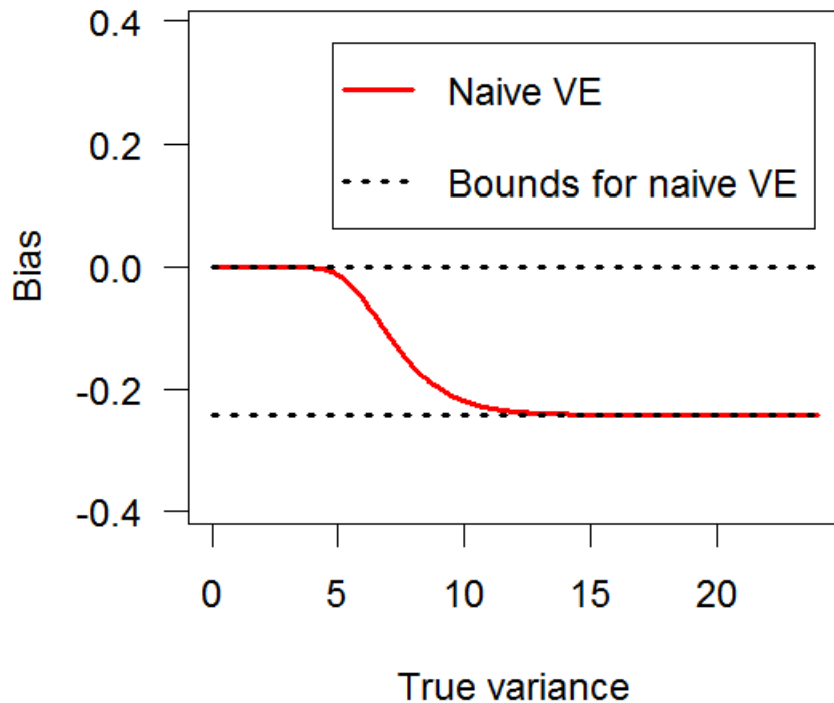


Bias of variance estimator, $n_{2\min}=0$

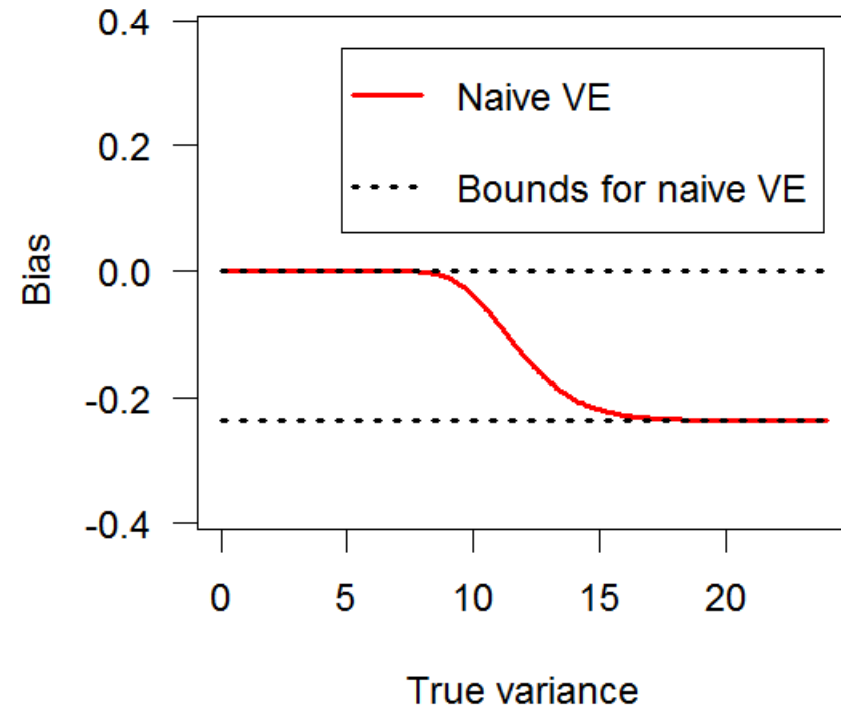


Bias for unblinded case with 1 IA

Bias of variance estimator, $n_1=20$

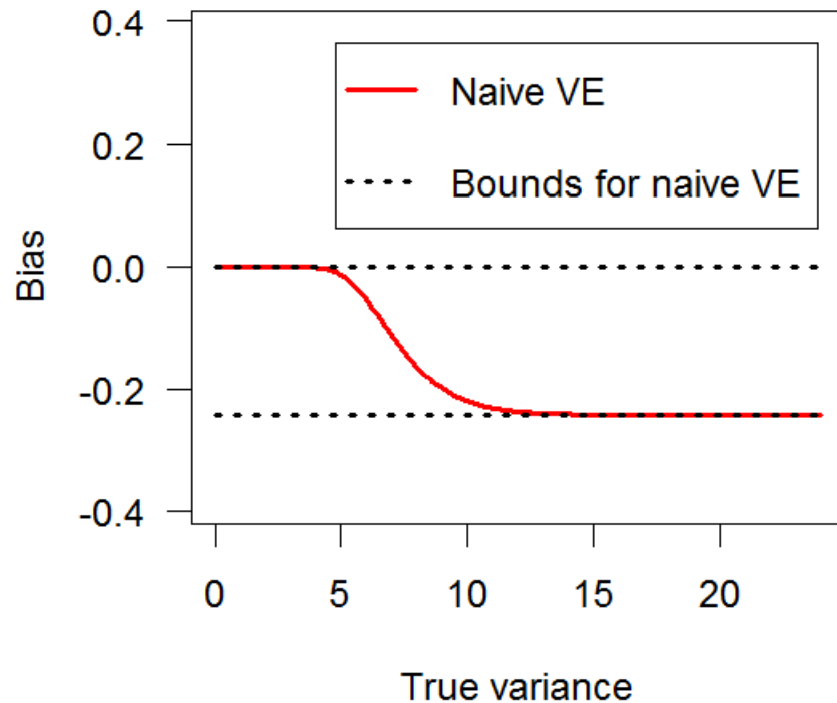


Bias of variance estimator, $n_1=40$

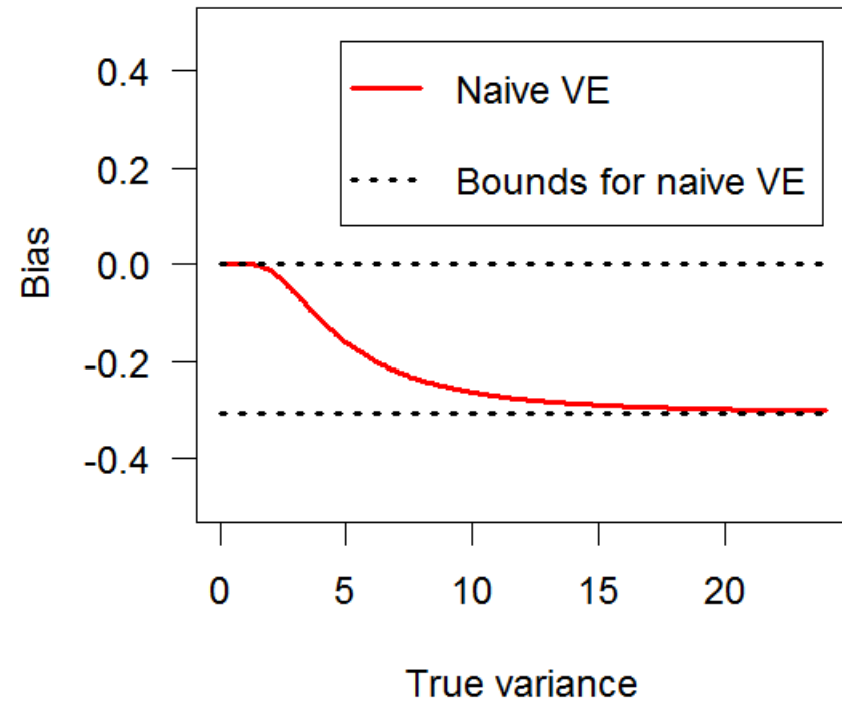


Bias for unblinded case with 1 IA

Bias of variance estimator, $n_1=20$



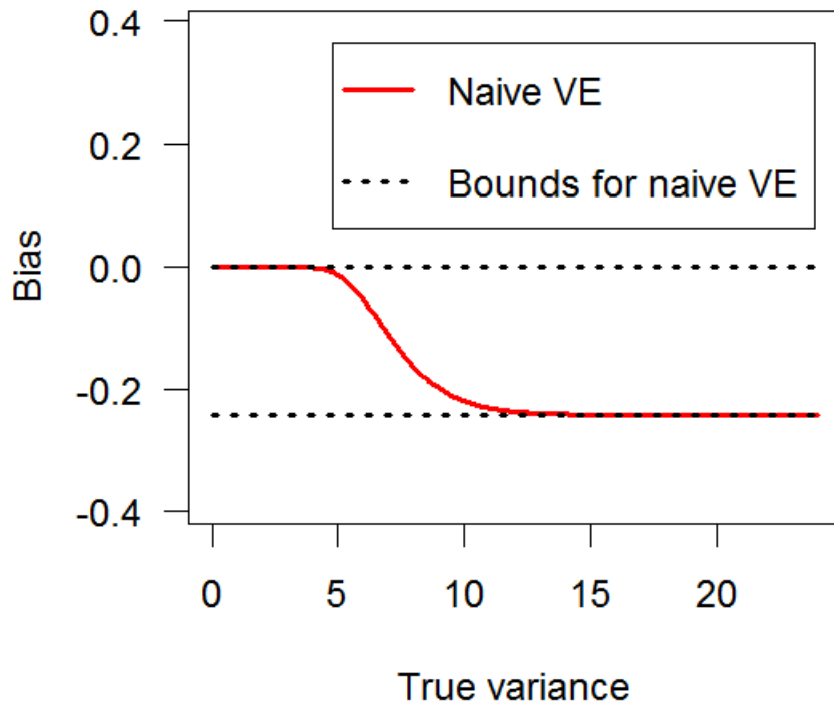
Bias of variance estimator, $n_1=5$



Bias for unblinded case with 1 IA

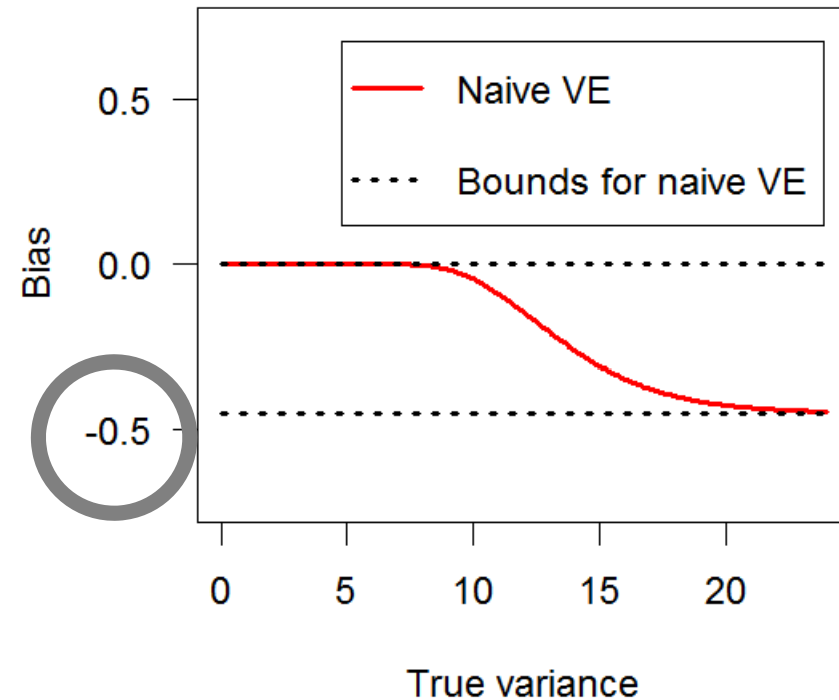
**Bias of variance estimator,
 $v=4.34$**

$(\alpha=0.025, \beta=0.1, \delta_a=2.2)$



**Bias of variance estimator,
 $v=2.34$**

$(\alpha=0.025, \beta=0.1, \delta_a=3)$



Maximal relative bias for both cases ~2.5%

Dependence of bias on parameters

- The **absolute bias** for large variances depends almost not on n_1 and not on the minimum number of patients in Stage 2
- The **relative bias** is largest if there is no minimum number of patients for Stage 2 and n_1 is small; however, decreasing n_1 to very small values (e.g. from 10 to 5) does not increase the relative bias much
- The relative bias in the considered scenarios was **at most 4%**

Summary

- Sample size re-estimation **can ensure appropriate power** (neither over- or under-powered) even under uncertainty of nuisance parameters
- The **effects on estimates and tests are usually small** and often, they might be totally acceptable
- In specific situations (small studies) it might be worth to investigate the bias further

References

- **Axén, Bodin, Bergström et al (2012)**. The use of weekly text messaging over 6 months was a feasible method for monitoring the clinical course of low back pain in patients seeking chiropractic care. *J. Clin. Epidem.* 65: 454-361.
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